

# Uniform, non-asymptotic confidence sequences for sequential estimation of average treatment effect

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Steve Howard

Joint work with Aaditya Ramdas, Jon McAuliffe, and Jasjeet Sekhon

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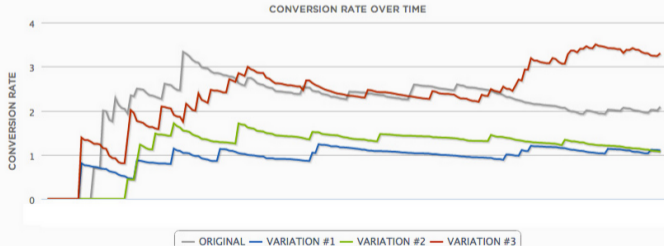
# Sequential monitoring of experiment results is problematic.

## Click on the button

The percentage of visitors who clicked on a tracked element.

✔ Variation #3 is beating Original by +58.0%.

VARIATIONS	VISITORS	CONVERSIONS	CONVERSION RATE	IMPROVEMENT	CHANCE TO BEAT BASELINE ?
Variation #3	970	32	3.3% ( $\pm 1.12\%$ )	+58.0%	95.2%
Original <small>BASELINE</small>	1,006	21	2.1% ( $\pm 0.88\%$ )	---	---
Variation #1	999	11	1.1% ( $\pm 0.65\%$ )	-47.3%	3.9%
Variation #2	1,027	11	1.1% ( $\pm 0.63\%$ )	-48.7%	3.3%



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- If the treatment effect is weaker than expected (or the budget has increased), we can extend the experiment.

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(Can compute always-valid p-values instead, if desired.)

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**Assumption:**  $Y_t(k) \in [0, 1]$  for  $k = 0, 1$ , all  $t$ .

- More on this later

## We define a sequence of average treatment effect estimands.

Our goal: after observing units  $1, \dots, t$ , we'd like to estimate

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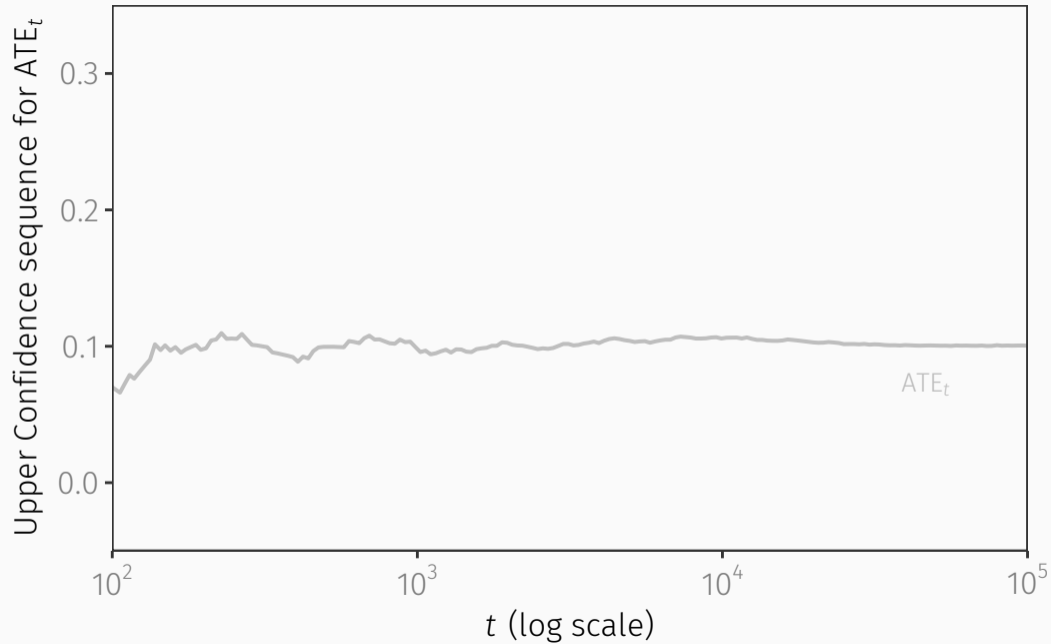
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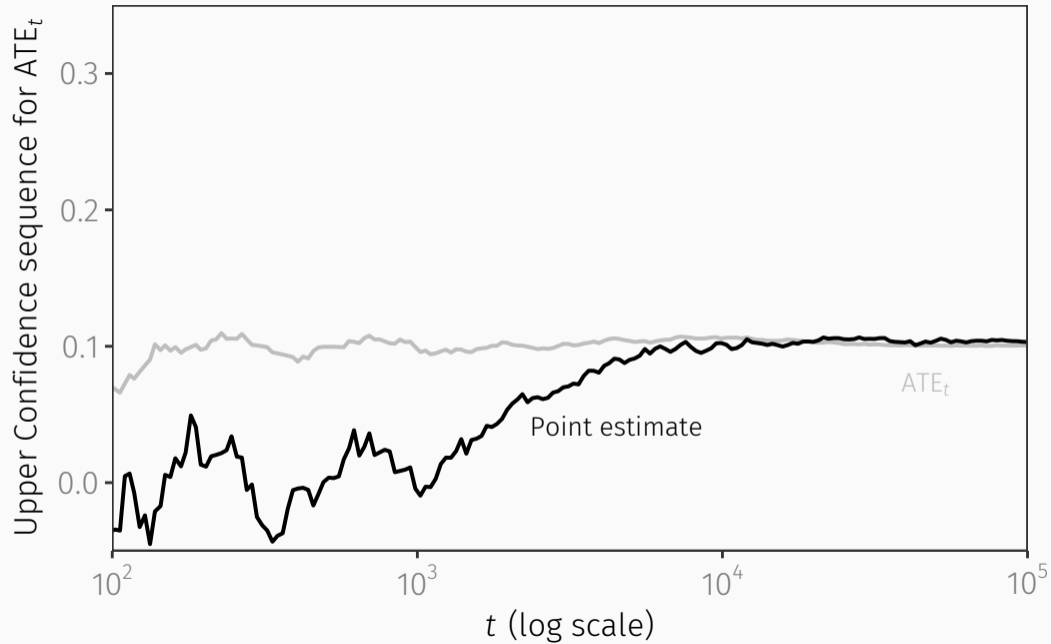
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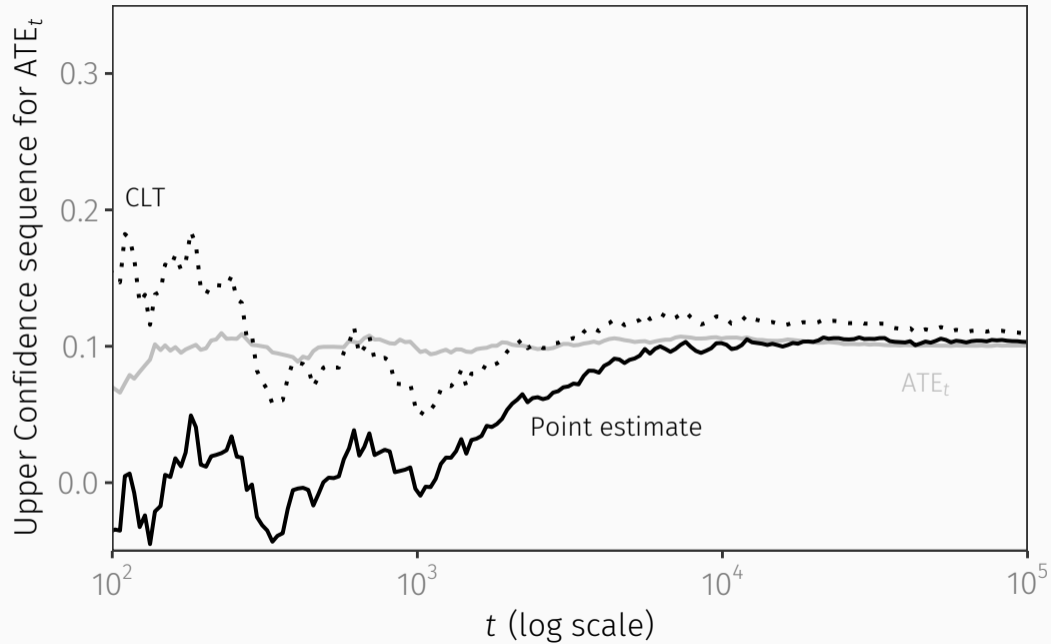
A **confidence sequence** for  $(\text{ATE}_t)_{t=1}^\infty$  is a sequence of intervals  $(\text{CI}_t)_{t=1}^\infty$  satisfying

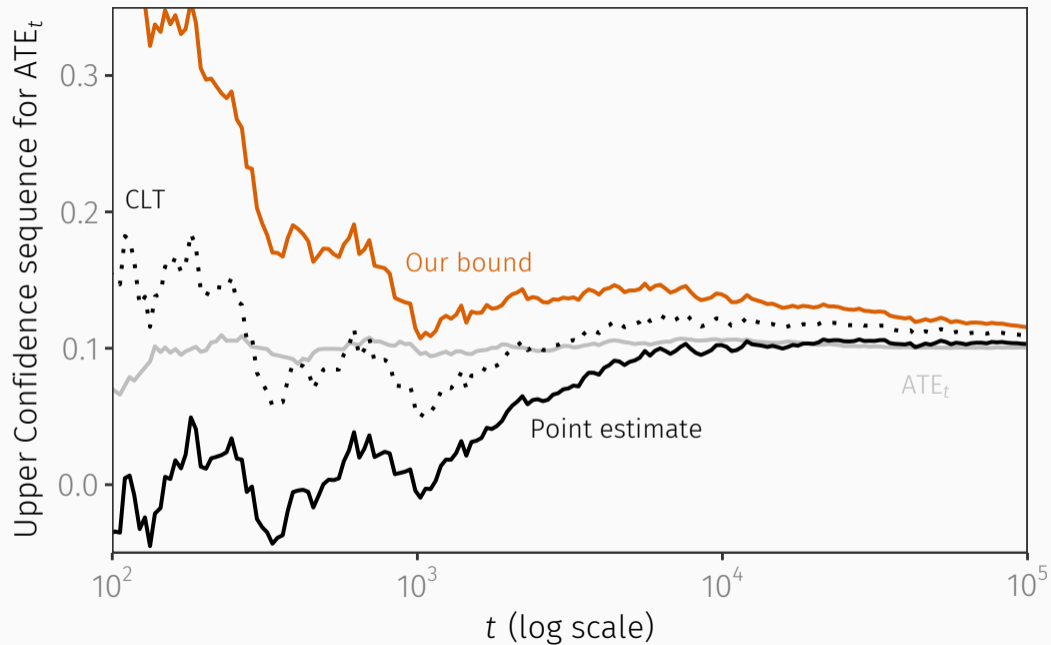
$$\mathbb{P}(\text{ATE}_t \in \text{CI}_t \text{ for all } t \in \mathbb{N}) \geq 1 - \alpha.$$

[Darling and Robbins 1967, Lai 1984, Jennison and Turnbull 1989, Johari et al. 2015]











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Two key properties of  $X_t$ :

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2. **Variance** of  $X_t$  depends on **prediction errors**  $(Y_t(1) - \widehat{Y}_t(1))^2$  and  $(Y_t(0) - \widehat{Y}_t(0))^2$ .

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Let  $S_t = \sum_{i=1}^t X_i$ . Then  $S_t/t$  is unbiased for  $ATE_t$ .

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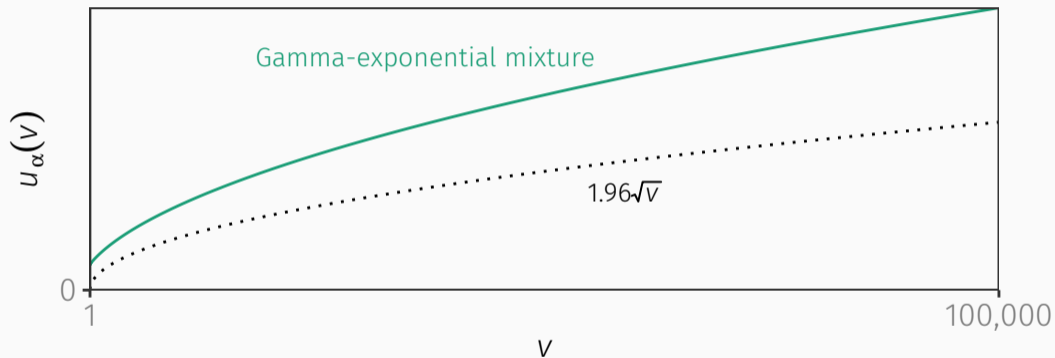
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- Estimation precision depends on prediction accuracy.
- $u_\alpha(v) = \mathcal{O}(\sqrt{v \log v})$ , so  $u_\alpha(v)$  is like  $z_{1-\alpha} \sqrt{v}$ , but the “z-factor” grows over time (slowly).



The uniform boundary grows only slightly faster than  $\mathcal{O}(\sqrt{n})$



## Theorem

Assuming no interference, if  $Y_t(k) \in [0, 1]$  for all  $k, t$ , then

$$\mathbb{P} \left( \left| \frac{S_t}{t} - \text{ATE}_t \right| < \frac{u_\alpha \left( \sum_{i=1}^t (X_i - \hat{X}_t)^2 \right)}{t} \text{ for all } t \in \mathbb{N} \right) \geq 1 - \alpha.$$

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This implies

$$\frac{S_t}{t} \pm \frac{u_\alpha \left( \sum_{i=1}^t (X_i - \hat{X}_i)^2 \right)}{t}$$

gives a  $(1 - \alpha)$ -confidence sequence for  $(\text{ATE}_t)$ .

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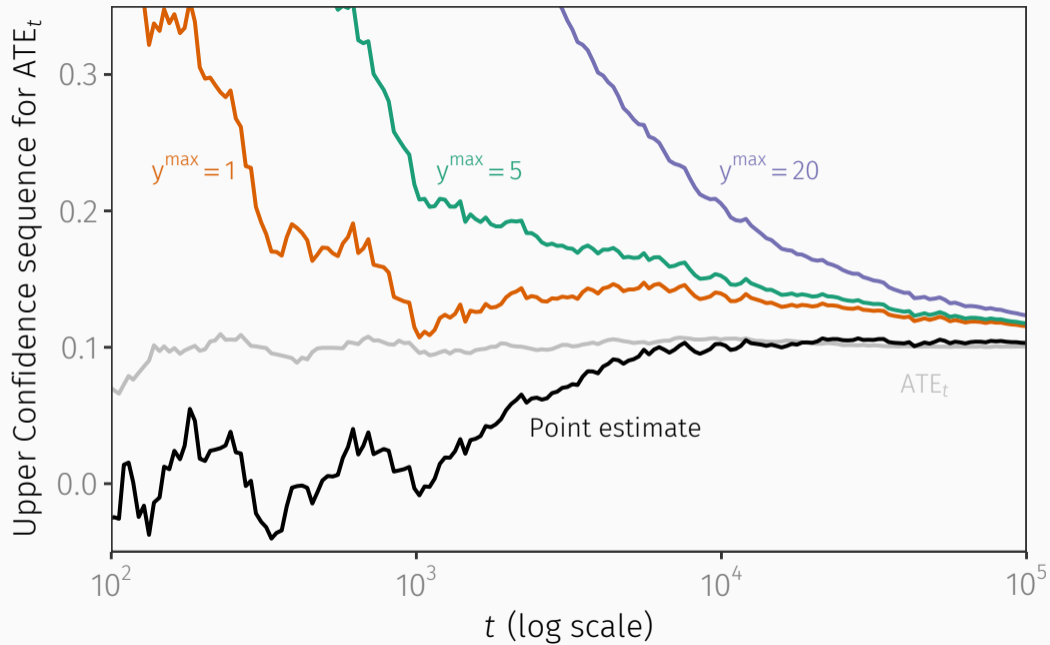
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Asymptotic arguments often sweep this issue under the rug.



## Recap

- Non-asymptotic confidence sequences for  $ATE_t$
- Flexible inferential tool for sequential experiments
- Provable coverage under the assumption of bounded potential outcomes
- Replace central limit theorem argument with uniform concentration bounds

Not covered today:

- Valid under superpopulation model as well
- Seamlessly handles biased coin designs or adaptive allocation
- Can estimate quantiles with similar tools (in i.i.d. framework)



Thank you!

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Uniform, nonparametric, non-asymptotic confidence sequences

`https://arxiv.org/abs/1810.08240`

Sequential estimation of quantiles

with applications to A/B-testing and best-arm identification

`https://arxiv.org/abs/1906.09712`